

Quau 203-2021

Qn ~~one~~ I.

The value of θ that maximizes $L(\theta)$ if it exists is some θ^* such that $L(\theta^*) \geq L(\theta)$.

Since \log is increasing,

$L(\theta^*) \geq L(\theta)$ this implies that $\log(\theta^*) = \log(\theta)$
i.e. $\ln(\theta)$.

Maximising $\ln(\theta)$ involves $\frac{d}{d\theta} (\ln(\theta))$ and setting it equal to zero and

Maximising $L(\theta)$ involves taking the derivative of $L(\theta)$ and setting it equal to zero

$$\frac{\partial \ln(\theta)}{\partial \theta} = \frac{L'(\theta)}{L(\theta)}$$

$$\text{Thus } \frac{\partial \ln(\theta)}{\partial \theta} = 0 = \frac{L'(\theta)}{L(\theta)} = 0$$

Thus $L'(\theta) = 0$ in both cases showing that their maximizer is the same.

Qn 2

$$(a) \int x^p (1-x) dx.$$

$$\int x^p - x^{p+1} dx.$$

$$= \frac{x^{p+1}}{p+1} - \frac{x^{p+2}}{p+2} + C$$

$$L.C.M = (p+1)(p+2)$$

~~constant of proportionality is (p+1)(p+2).~~

$$\frac{(p+2)x^{p+1} - (p+1)x^{p+2}}{(p+1)(p+2)}$$

constant of proportionality is (p+1)(p+2)

$$(b) \int x(x^p (1-x)) dx.$$

$$\int x^{p+1} (1-x) dx.$$

$$\int x^{p+1} - \int x^{p+2} dx$$

$$\frac{x^{p+2}}{p+2} - \frac{x^{p+3}}{p+3} = \frac{p+1}{p+3}$$

$$(c) \ln = \prod_{i=1}^n x^p (1-x).$$

$$\ln = \prod_{i=1}^n (x^{p+1} - x^{p+2})$$
$$\prod_{i=1}^n (p+1) x_p^{\wedge}$$

introducing logarithms.

$$\ln(L_p) = n \ln(p+1) + n \ln(p+2) + p \sum \ln x_i + \sum \ln(1-x_i)$$

Qn 3

$$(a) f(x) = \theta x^{\theta-1}, \quad 0 < x < 1, \quad \theta > 0$$

$$L(\theta) = \prod_{i=1}^n \theta x_i^{\theta-1}$$

factor θ we have

$$l(\theta) = \hat{\theta} \left(\prod_{i=1}^n x_i \right)^{\theta-1}$$

introducing natural log

$$\begin{aligned} \ln(L(\theta)) &= \ln \theta^n \left(\prod_{i=1}^n x_i \right)^{\theta-1} \\ &= n \ln \theta + (\theta-1) \ln \left(\prod_{i=1}^n x_i \right) \end{aligned}$$

$$= n \ln \theta + (\theta-1) \sum \ln x_i$$

MLE is calculated by taking derivative of $\theta \ln(L)(\theta)$.

$$\left[n \ln(\theta) + (\theta-1) \ln \left(\prod_{i=1}^n x_i \right) \right]' = \frac{n}{\theta} + \ln \left(\prod_{i=1}^n x_i \right)$$

$$\Rightarrow \theta_0 = \frac{-n}{\ln \left(\prod_{i=1}^n x_i \right)} = \frac{-1}{\ln(x)} \quad \text{is the MLE of } \theta$$

$$(b) f(x) = \theta x^{\theta-1}, \quad 0 < x < 1$$

$$E(x) = \int_0^1 x \cdot f(x) dx$$

$$= \int_0^1 x \cdot \theta x^{\theta-1} dx$$

$$= \theta \int_0^1 x \cdot x^{\theta-1} dx$$

$$= \theta \int_0^1 x dx$$

$$\theta \int_0^1 x^{\theta} dx$$

$$\theta \left[\frac{x^{\theta+1}}{\theta+1} \right]_0^1 = \theta \left[\frac{1}{\theta+1} \right] =$$

$$\frac{\theta}{\theta+1}$$

(c) ML \neq MM i.e. Maximum likelihood estimator is not equal to method's of moment estimator.

(4)
 (a) $f(x) = \frac{\lambda^{\nu}}{\Gamma(\nu)} x^{\nu-1} \exp(-\lambda x)$

$$\ln(\nu) = \prod_{i=1}^n \frac{\lambda^{\nu}}{\Gamma(\nu)} x_i^{\nu-1} \exp(-\lambda x_i)$$

$$= \left(\frac{\lambda^{\nu}}{\Gamma(\nu)} \right)^n \prod_{i=1}^n x_i^{\nu-1} \exp(-\lambda x_i)$$

$$= \left(\frac{\lambda^{\nu}}{\Gamma(\nu)} \right)^n \left(\prod_{i=1}^n x_i \right)^{\nu-1} \exp(-\lambda \sum_{i=1}^n x_i)$$

Dropping the term that does not contain ν

The log likelihood is

$$\ln(\nu) = n\nu \log \lambda - n \log \Gamma(\nu) + (\nu-1) \log \left(\prod_{j=1}^n x_j \right)$$

(b) $M(t) = \left(\frac{\lambda}{\lambda+t} \right)^{\nu}$

$$\text{Mean} = \frac{d}{dt} M(t) = \frac{d}{dt} \left(\frac{\lambda}{\lambda+t} \right)^{\nu} = \nu (\lambda(\lambda+t)^{-1})^{\nu-1} (-1)(\lambda+t)^{-1} \Big|_{t=0}$$

$$= \nu (\lambda(\lambda+t)^{-1})^{\nu-1} (-1)(\lambda+t)^{-1} \Big|_{t=0} = -\nu$$

$$\frac{d^2}{dt^2} M(t) = \frac{d}{dt} \left(\nu (\lambda(\lambda+t)^{-1})^{\nu-1} (-1)(\lambda+t)^{-1} \right)$$

=

$$\gamma(\gamma-1)(\lambda(\lambda-t)^{-1})^{\gamma-2} (\lambda+(\lambda-t)^{-1}(\lambda+(\lambda-t)^{-1})) +$$

$$(1-(\lambda-t)^{-2})(\gamma(\lambda(\lambda-t)^{-1})^{\gamma-1})/t_0$$

$$= \left(\gamma - \frac{\gamma}{\lambda^2}\right) (\gamma^2 - \gamma) = \gamma^3 - \gamma^2 - \frac{\gamma^3}{\lambda^2} + \frac{\gamma^2}{\lambda^2} -$$

$$\begin{aligned} \text{Variance} &= \gamma^3 - \gamma^2 - \frac{\gamma^3}{\lambda^2} + \frac{\gamma^2}{\lambda^2} - (\gamma)^2 \\ &= \gamma^3 - 2\gamma^2 - \gamma^3 - \left(\frac{\gamma^3 - \gamma^2}{\lambda^2}\right) \end{aligned}$$

The maximum likelihood estimators are not equal (\neq).